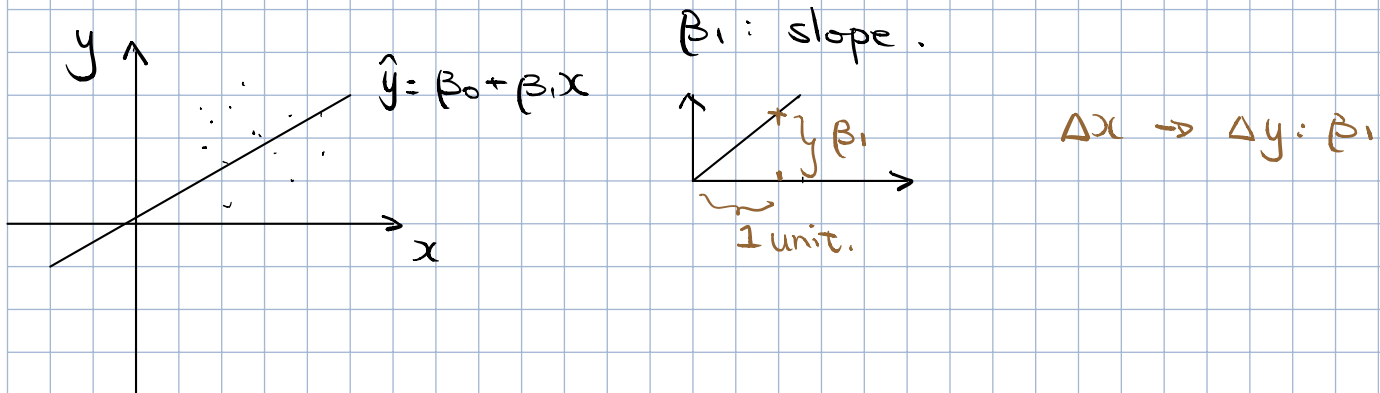


Intuition on Stats.

Start from a simple linear regression, which has 1 predictor.



model: $\hat{y} = \beta_0 + \beta_1 x$.

model assumption: we expect / model predicts

\hat{y} : on average

slope: $\Delta x \rightarrow \Delta y: \beta_1$

"1" Δx : - For numeric var, 1 unit.

- For categorical var with k levels,

"1" can represent 1 2 3 ... k

any gap between two neighbors

"distance" between neighbors may vary!

$|k-1|$ different representation of "1" Δx

need $|k-1|$ different β 's to estimate Δy resulted from different "1" in Δx .

method 1. 1 2 3 4 ... k \Rightarrow baseline change.

method 2. 1 2 3 4 ... k \Rightarrow baseline fix:
level 1.

$$1 \Delta X \rightarrow \Delta y = \begin{matrix} \beta_0 \\ \beta_0 + \beta_{1,1} \\ \beta_0 + \beta_{1,2} \\ \vdots \\ \beta_0 + \beta_{1,k-1} \end{matrix} \begin{matrix} \text{level 1} \\ \text{level 2} \\ \text{level 3} \\ \vdots \\ \text{level k} \end{matrix}$$

baseline increment

Interpret $\beta_{1,1}, \dots, \beta_{1,k-1}$ as "increment":

- "slope": $\Delta X \rightarrow \Delta y$

Compare to level 1 (baseline), level 2 has $\Delta y: \beta_{1,1}$

model: $\hat{y} = \beta_0 + \beta_{1,1} I(\text{level 2}) + \beta_{1,2} I(\text{level 3}) + \dots + \beta_{1,k-1} I(\text{level k})$

model assumption: we expect / model predicts

\hat{y} : on average

slope: $\Delta X \rightarrow \Delta y: \beta_1$

compared to level 1, level 2 will have $\Delta y: \beta_1$

For model with 1 predictor, (categorical)

$$\hat{y} = \beta_0 + \beta_{1,1} I(\text{level 2}) + \beta_{1,2} I(\text{level 3}) + \dots + \beta_{1,k-1} I(\text{level k})$$

	I(level 2)	I(3)	I(4)	...	I(k)
level 1	0	0	0		0
2	1	0	0		0
3	0	1	0		0
...					
k	0	0	0		1

Already multivariate linear regression !

Special case ,

one observation can only have 1 level ,

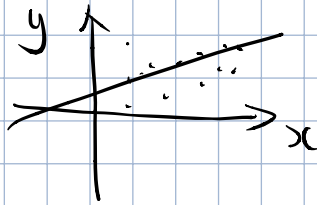
SLR.

y, x



EDA

→ linear trend.



← linear regression

Assumption: $y = \beta_0 + \beta_1 x + \epsilon$

model fitting: $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$

interpretation of $\hat{\beta}_0, \hat{\beta}_1$

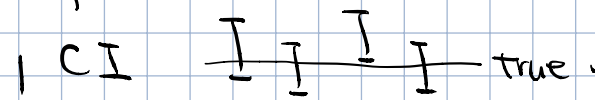
prediction : $\hat{y}^* = \hat{\beta}_0 + \hat{\beta}_1 x^*$

extrapolation.

evaluation : R^2

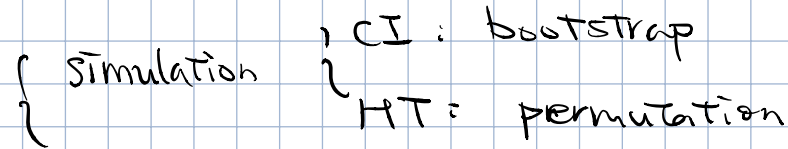
RMSE.

Inference :



$\beta_0 = 0$ vs $\beta_0 \neq 0$

method



math. ($\epsilon \stackrel{iid}{\sim} N(0, \sigma_\epsilon^2)$)

CI $\hat{\beta} \pm t^* \times \text{se}(\hat{\beta})$ → degree of freedom

HT : $t^* = \frac{\hat{\beta}}{\text{se}(\hat{\beta})}$ p-val = $P_r(|t| > t^*)$

Interpretation.

Predictive Interval.

Diagnosis:

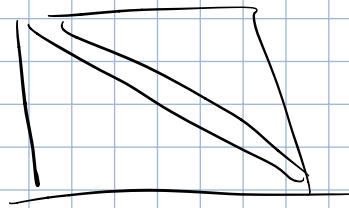
Condition. (F) ← residual plot

Influential pts:

leverage
standardized residuals
Cook's Distance

MLR.

EDA : ggpair



Model assumption : $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p + \epsilon$.

Model fitting : $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p$

Interpretation : - Hold constant - on average

prediction : - expect - $\Delta x \rightarrow \Delta y$

$$\hat{y}^* = f(x^*)$$

Type of predictors :

- mean-center

- categorical : dummy variable

- Interaction : $\Delta \beta$

- Transformation : $\log(\cdot)$; $x^2 + x$

Model comparison :

ANOVA

Adj R^2

AIC

BIC

Feature engineering -

split data

recipe .

workflow

fit model

prediction .

CV

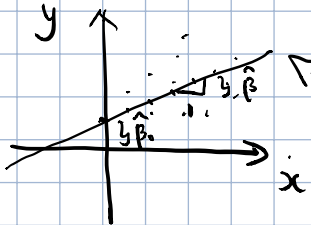
Conditions : residual plot

↑
multicollinearity .

Simple Linear Regression.

$y \sim x$

EDA.



Model Assumption: $y = \beta_0 + \beta_1 x + \epsilon$

Fit: $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x \rightarrow$

tidy(): point. se. t-stat. p-val.

meaningful? $\rightarrow \hat{\beta}_0 = x=0 \Rightarrow \hat{y} = \hat{\beta}_0$
 center x. $\hat{\beta}_1: \Delta x \rightarrow \Delta \hat{y}: \hat{\beta}_1$
 interpret.

- * we expect / predict.
- * on average
- * 1 unit. $\rightarrow \hat{\beta}_1$

prediction: $y^* = \hat{\beta}_0 + \hat{\beta}_1 x^*$

$x^* \rightarrow$ extrapolation.

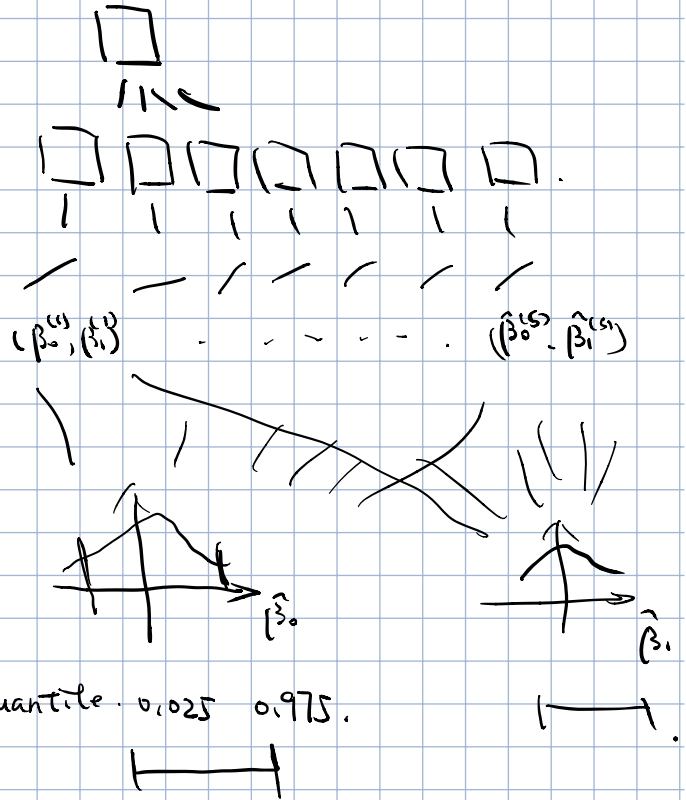
evaluation: R^2 \uparrow
 RMSE. \downarrow

Inference. $\left\{ \begin{array}{l} CI \\ HT \end{array} \right.$ $\overline{I I I I I}$

Interpret. CI: - (95%)

- model predict / expect.
- on average.
- $\Delta x \uparrow$ [,]
- CI for β
- CI for \bar{y} \rightarrow individual.
- prediction interval. $>$ CI. \bar{y} \rightarrow average mean.

Simulation. Bootstrap: sample w/ replacement
 math.



math: $y = \beta_0 + \beta_1 x + \epsilon$

$\epsilon \sim N(0, \sigma_\epsilon^2)$

$\hat{\beta} \pm t_{df}^* \times \text{se}(\hat{\beta})$

$df = n - p - 1 = n - 2$
 # predictor

HT:

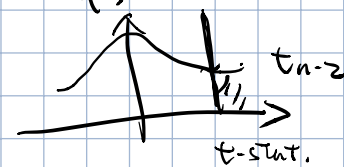
$H_0: \beta_1 = 0$ vs: $H_1: \beta_1 \neq 0$

Assume H_0 is true.

check obs. is reasonable.

Math: t-stat. $\frac{\hat{\beta}_1}{\text{se}(\hat{\beta}_1)} \sim t_{n-2}$

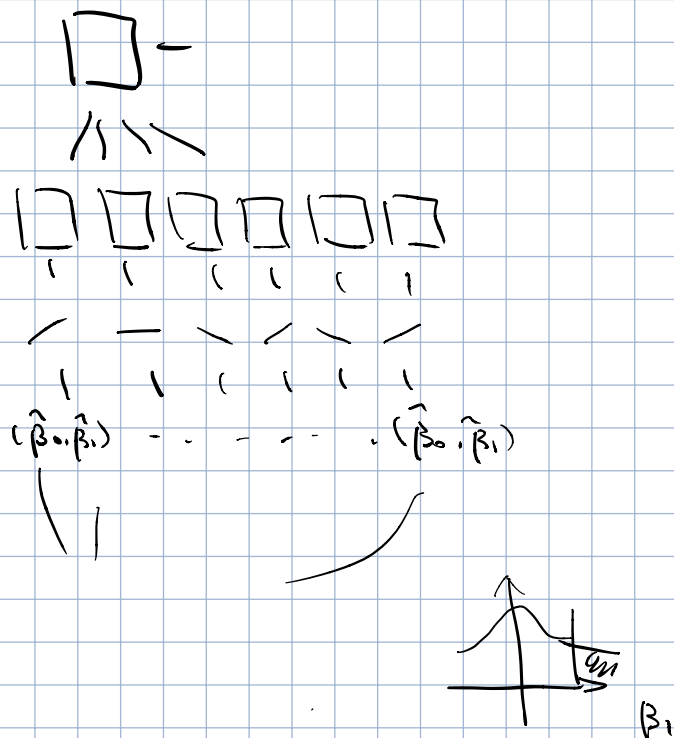
p-val



$\Pr(|T_i^*| > t)$

p-val. small.
 reject H_0 .

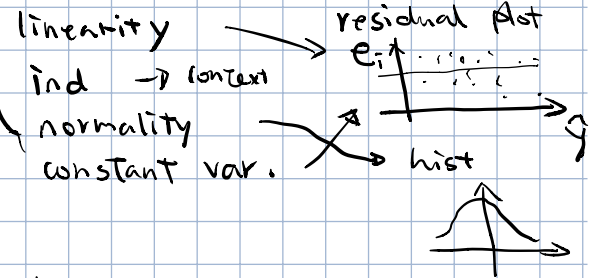
Simulation: permutation. (sample w/o replacement)



calculate p-val.

Diagnosis:

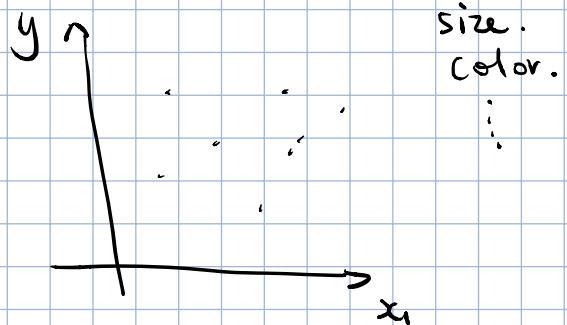
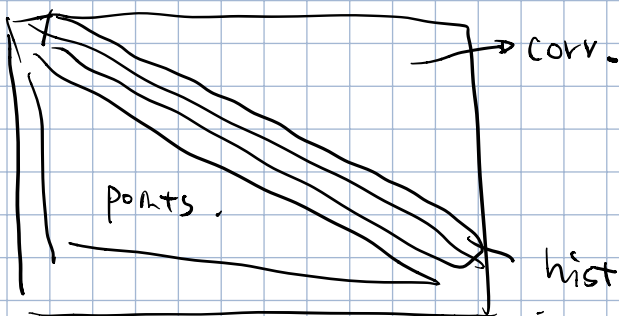
- check conditions.
- influential points.



- outlier. ← standardize res. > 3
- high leverage. ←
- Cook dist. : 0.5, 1.

MLR.

EDA: ggpair. (.)



Model : $y = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p + \epsilon$

Fit : $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p + \epsilon$

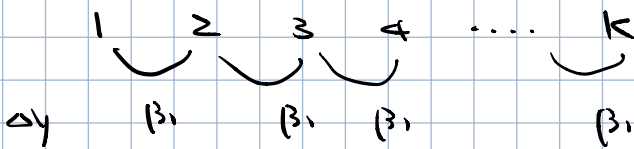
hold other constant.

Interpretation. $\hat{\beta}_1$ (4) - on avg - we expect. - 1 unit $\rightarrow \dots \Delta y : \beta_1$
 - hold other constant.

Δx : numeric var $\Delta x \rightarrow \Delta y : \hat{\beta}_1$

Δx : 1 unit increase.

Δx : categorical var. k level. Δ



Δy 1 2 3 4 ... k fix reference level 1.

k levels.

$\beta_{1,1} \beta_{1,2} \dots \beta_{1,k-1}$

(k-1) dummy.

$I(2) \quad I(3) \quad \dots \quad I(k)$

1	0	0	...	0
2	1	0	...	0
...				
k	0	0	...	1

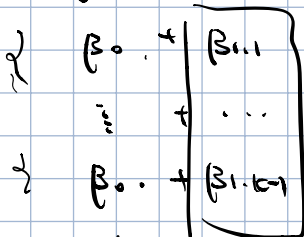
$y = \beta_0 + \beta_1 x_1 + \dots$



x_1 is k-level cate...

$y = \beta_0 + \beta_{1,1} I(x_1=2) + \dots + \beta_{1,k-1} I(x_1=k)$

β_0 increment. - $\Delta x \rightarrow \Delta y$: compare to level 1.

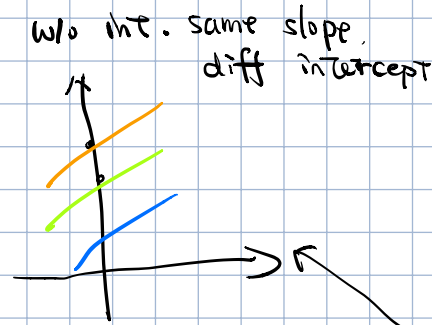
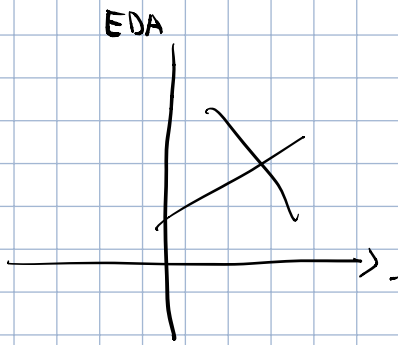


we expect the y of obs w/ level 2. will increase by $\beta_{1,1}$, on avg. hold other constant.

$\hat{\beta}_0$: obs. with level ...
 baseline.

Interpretation : 1. numeric vs. categorical.

2. interaction.



$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_1 X_2$$

X_1 : numeric X_2 : categorical.

w/o

$$Y = (\beta_0 + \beta_2 X_2) + \beta_1 X_1$$

↑
intercept.

$$Y = (\underbrace{\beta_0 + \beta_2 X_2}_{\text{inter}}) + (\underbrace{\beta_1 + \beta_3 X_2}_{\text{slope}}) X_1$$

$$\begin{matrix} \beta_0 \\ \beta_0 + \beta_{2,1} \\ \vdots \\ \beta_0 + \beta_{2,k-1} \end{matrix}$$

baseline + increment.

$$\begin{matrix} \beta_1 \\ \vdots \\ \beta_1 + \beta_{3,1} \\ \vdots \\ \beta_1 + \beta_{3,k-1} \end{matrix}$$

↓

Δy

$\Delta \beta$:

the potential influence of

$\hat{\beta}_{3,1}$ - $\Delta X_2 \Rightarrow \Delta \beta$: $\hat{\beta}_{3,1}$
 - expect.
 - on avg
 - hold other const.

Compare to level 1, we expect, obs w/ level 2, on outcome, increase by $\hat{\beta}_{3,1}$ on avg, hold other constant.

- mean center.

$$\Delta X \rightarrow \Delta y$$

- log. transformation.

$$y = \beta_0 + \beta_1 \log X_1$$

$$\log(X_1 + \Delta X)$$

$$\log(X_1 \cdot c) = \log X_1 + \log c$$

$$\Delta X \rightarrow \Delta y.$$

$$Y = \beta_0 + \beta_1 \log(1.01X) = \beta_0 + \beta_1 \log X + \beta_1 \log 1.01$$

X increase by 1%.

$$y. \beta_1 \log 1.01$$

$$y = \beta_0 + \beta_1 \log x + \epsilon \quad ? \text{ linear regression.}$$

Yes.

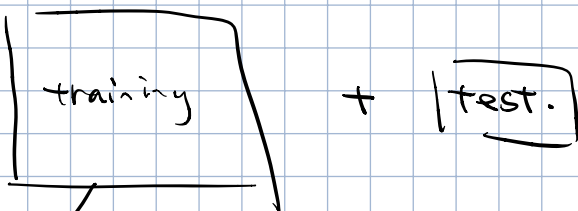
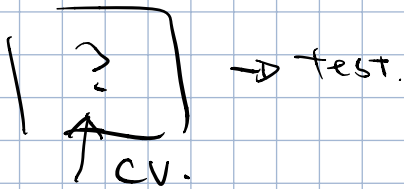
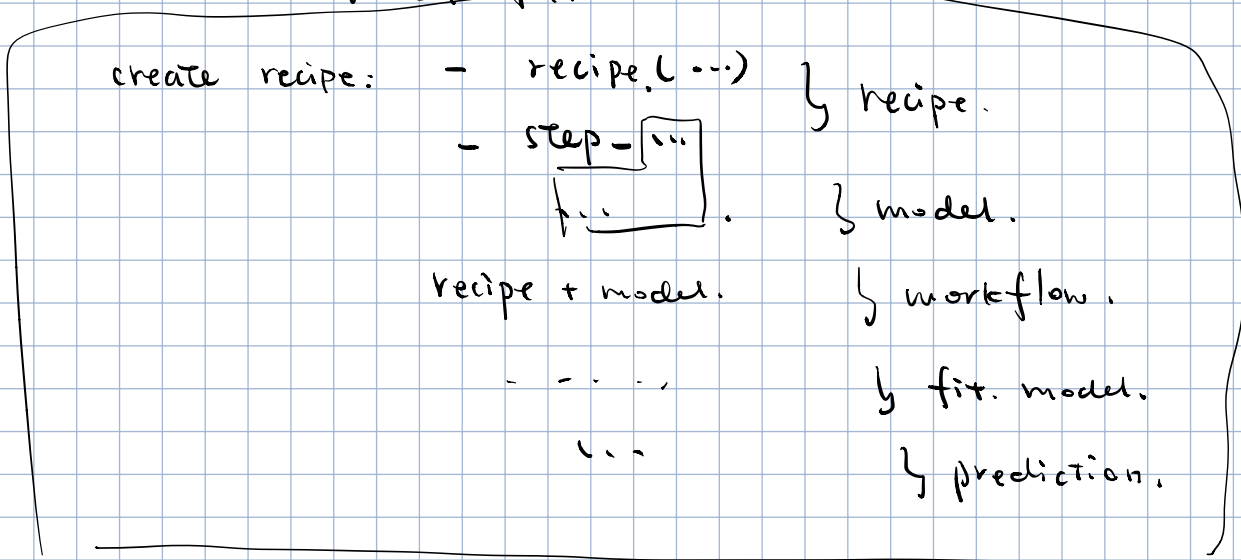
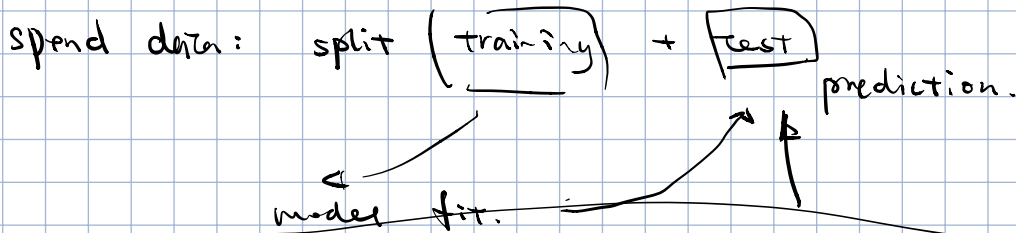
linearity wrt. β .

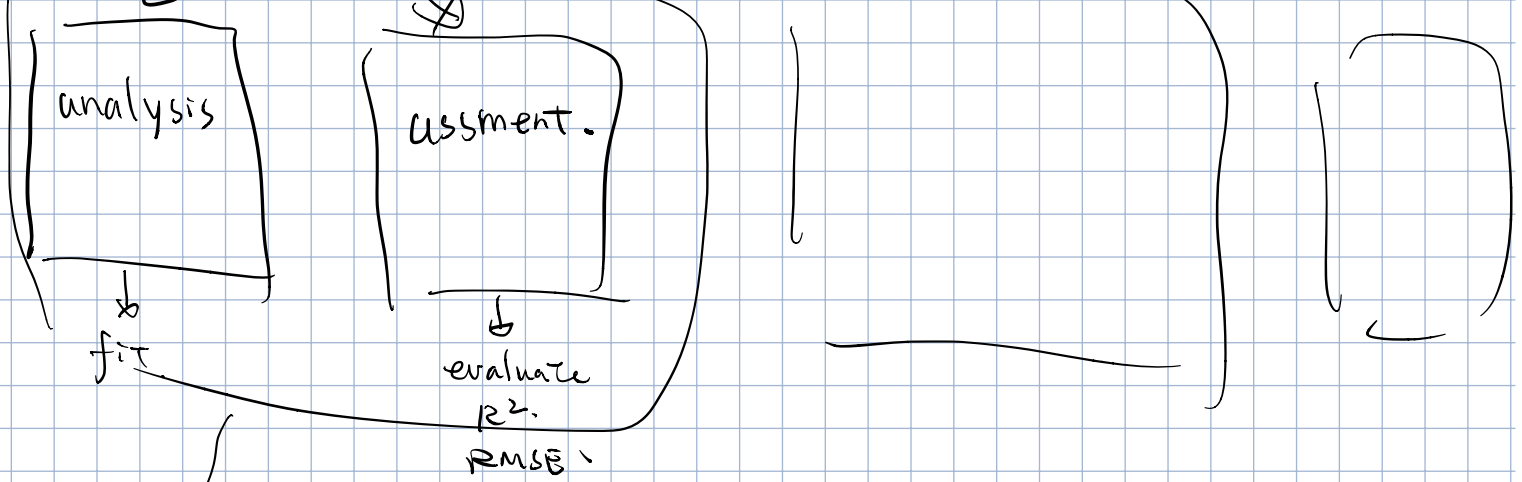
Type of predictors.

Model comparison:

- anova. $H_0: \beta_1 = \dots = \beta_k = 0$.
- p-val. small. reject H_0 .
- Adj. R^2 : \uparrow
- AIC, BIC: \downarrow .

Feature engineer: extract feature from predictor.
represent predictor.

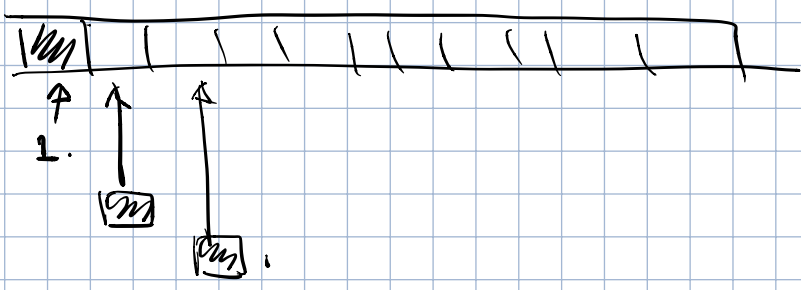




CV.

RMSE
R ²

10-fold



CV. } ev. fold. ()
 fit. samples ()

Inference. | CI.

MT. = $H_0: \beta_1 = 0$. hold other constant

Diagnosis : } conditions : (4). residual plot. - SLR.
 } multicollinearity, → non-identifiable.

↓
 se (-) ↑
 ↓
 inference imprecise.

detect: { corr. > 0.9.
 abnormal. coeff.
 VIF. : > 10. }

Select one to remove.



model comparison

