



Modeling Neural Population Coordination via a Block Correlation Matrix

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Research Question

- Estimation of a **block correlation matrix**
- **Unknown block structure**: grouping w.r.t. variables
- **Flexibility**: off-diagonal correlation $\in (-1,1)$
- **Interpretability**: model assumptions + priors
- **Statistical efficiency**: large p small n cases
- **Computational efficiency**: conjugate priors

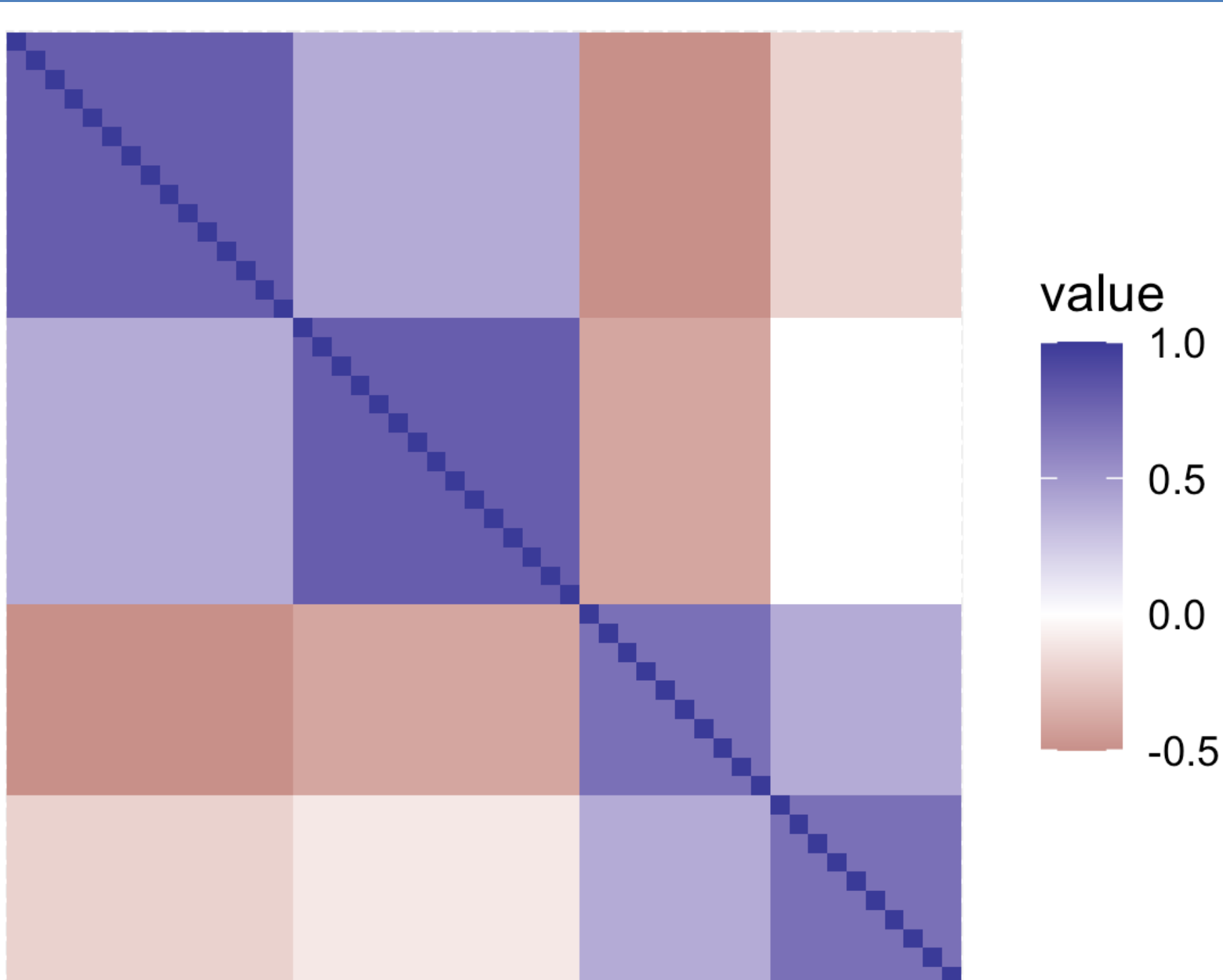
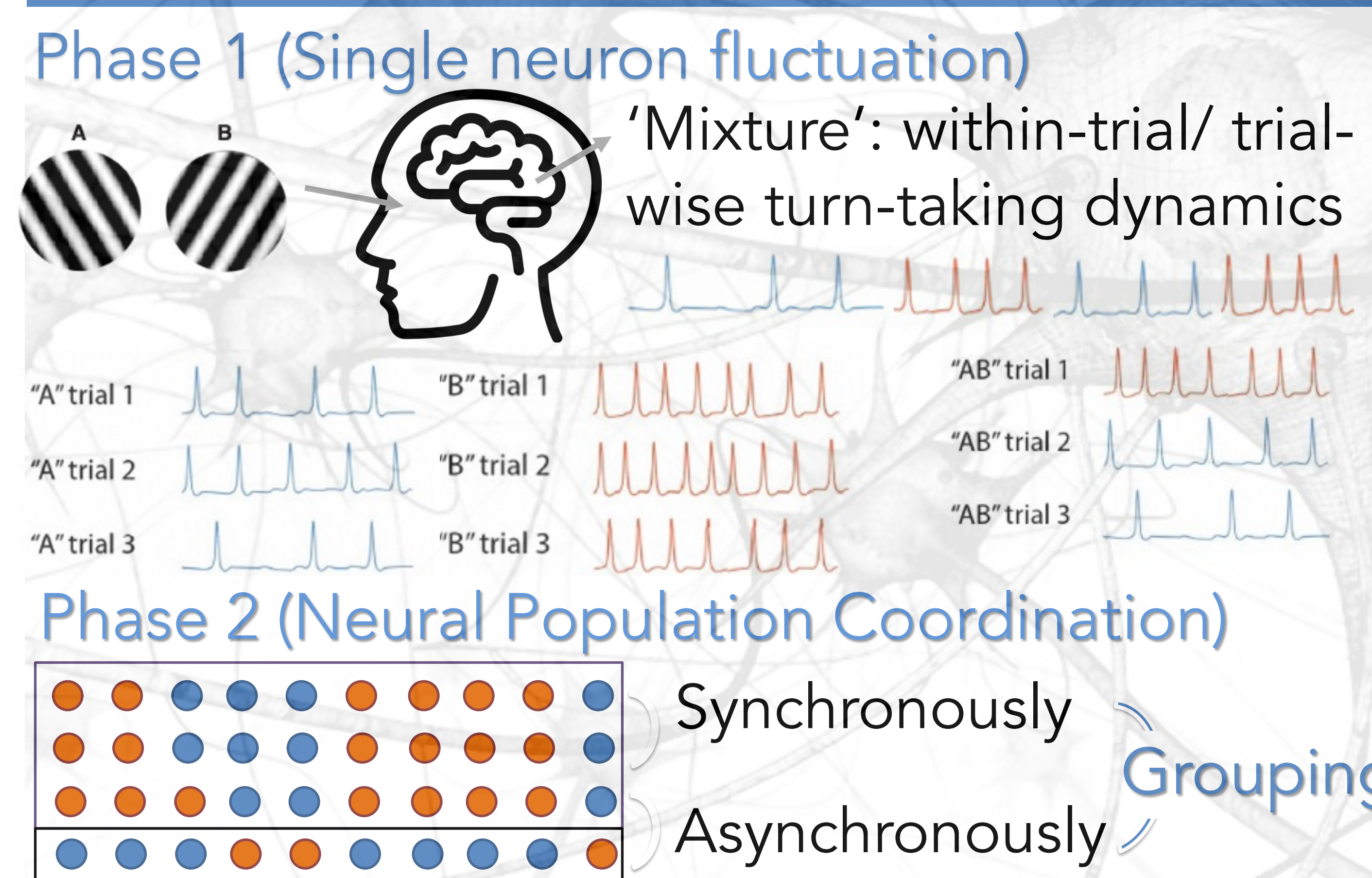


Fig. 1: An example of a block correlation matrix (50 variables 4 blocks)

Motivation



Method: Bayesian Block Correlation Matrix Estimation

Bayesian Model: Mixture of Finite Mixtures + Canonical Representation

Groups allocation: Mixture of Finite Mixtures (MFM) (Miller and Harrison, 2018)

\mathcal{C} denotes partition of $[N]$ induced by S_1, \dots, S_N :

$K \sim p_K$, where p_K is a p.m.f on $\{1, 2, \dots\}$, where we consider $K - 1 \sim \text{Pois}(1)$

Group allocation $S_i \in \{1, \dots, K\}$ $(\pi_1, \dots, \pi_k) \sim \text{Dir}_k(\gamma, \dots, \gamma)$ given $K = k$

$S_1, \dots, S_N \sim \pi$ (iid) given π

Permuted Data:

$$\tilde{Y} = PY, \quad \tilde{Y} \sim N(0, \Sigma),$$

Block Covariance Matrix:

$$\Sigma = \begin{pmatrix} \Sigma_{[1,1]} & \Sigma_{[1,2]} & \dots & \Sigma_{[1,K]} \\ \Sigma_{[2,1]} & \Sigma_{[2,2]} & \dots & \Sigma_{[2,K]} \\ \vdots & \vdots & \ddots & \vdots \\ \Sigma_{[K,1]} & \Sigma_{[K,2]} & \dots & \Sigma_{[K,K]} \end{pmatrix}, \quad \Sigma_{[i,i]} = \begin{pmatrix} \sigma_{ii}^2 & \sigma_{ii} & \dots & \sigma_{ii} \\ \sigma_{ii} & \sigma_{ii}^2 & \dots & \sigma_{ii} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{ii} & \sigma_{ii} & \dots & \sigma_{ii}^2 \end{pmatrix}, \quad \Sigma_{[i,j]} = \begin{pmatrix} \sigma_{ij} & \dots & \sigma_{ij} \\ \vdots & \ddots & \vdots \\ \sigma_{ij} & \dots & \sigma_{ij} \end{pmatrix}$$

Block Correlation Matrix maintain SAME block structure if: $\sigma_i^2 \neq \sigma_{ii}$

Canonical Representation: $\Sigma = QDQ'$,

(Archakov and Hansen, 2020)

Representatives + Replicates

$$D = \begin{pmatrix} A & 0 & \dots & 0 \\ 0 & \lambda_1 I_{n_1-1} & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & \lambda_K I_{n_K-1} \end{pmatrix}, \quad Q = \begin{pmatrix} v_{n_1} & 0 & \dots & v_{n_1\perp} & 0 & \dots & 0 \\ 0 & v_{n_2} & \dots & 0 & v_{n_2\perp} & \dots & \vdots \\ \vdots & \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \dots & v_{n_K} & 0 & \dots & \dots & v_{n_K\perp} \end{pmatrix}$$

$$a_{ij} = \sigma_{ij} \sqrt{n_i n_j}, \quad a_{ii} = \sigma_i^2 + (n_i - 1) \sigma_{ii} \quad v_{n_k} = \left(\frac{1}{\sqrt{n_k}}, \dots, \frac{1}{\sqrt{n_k}} \right)' \text{ Group size (Gram-Schmidt)}$$

$$\lambda_i = \sigma_i^2 - \sigma_{ii}, \quad QQ' = Q'Q = I.$$

$$\text{Log likelihood: } -\frac{NJ}{2} \log 2\pi - \frac{J}{2} \log |A| - \frac{1}{2} \sum_{j=1}^J \eta_{j(0)}^{*'} A^{-1} \eta_{j(0)}^* - \frac{J}{2} \sum_{k=1}^K (n_k - 1) \log \lambda_k - \frac{1}{2} \sum_{j=1}^J \sum_{k=1}^K \eta_{j(k)}^{*'} \eta_{j(k)}^* / \lambda_k,$$

where $\eta_j^* = Q' \tilde{Y}_j$,

Conjugate Priors:

$$A | \mathcal{C} \sim IW(\nu, A_0),$$

$$\lambda_k \sim (\text{iid}) IG(a_\lambda, b_\lambda).$$

Gibbs Sampler

- Initialize with one group.
- Within each iteration,
 - (1) For each variable, update its group allocation based on MFM's algorithm (adaptation of 'Algorithm 3' (Neal, 2000)).
 - (2) Update parameters by conjugacy.

Prior Specification

- Non-informative priors require:
 - (1) invariant to group size
 - (2) between-group: uniformly distributed
 - (3) within-group: positive, relatively high
- Scale with group size: $\nu = K, A_{0(kk)} = n_k, \lambda_k^{-1} \sim Ga(20, 10)$

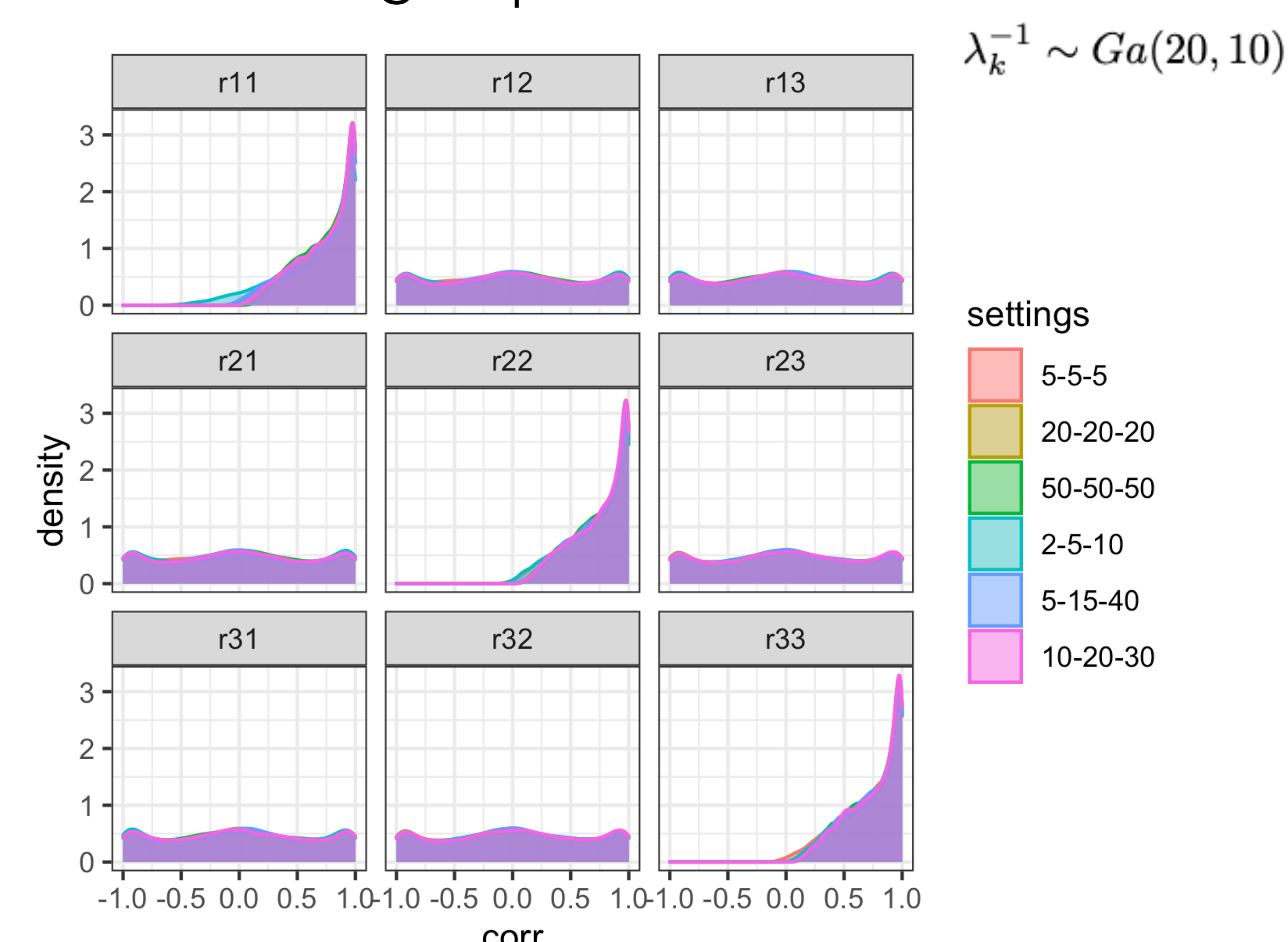
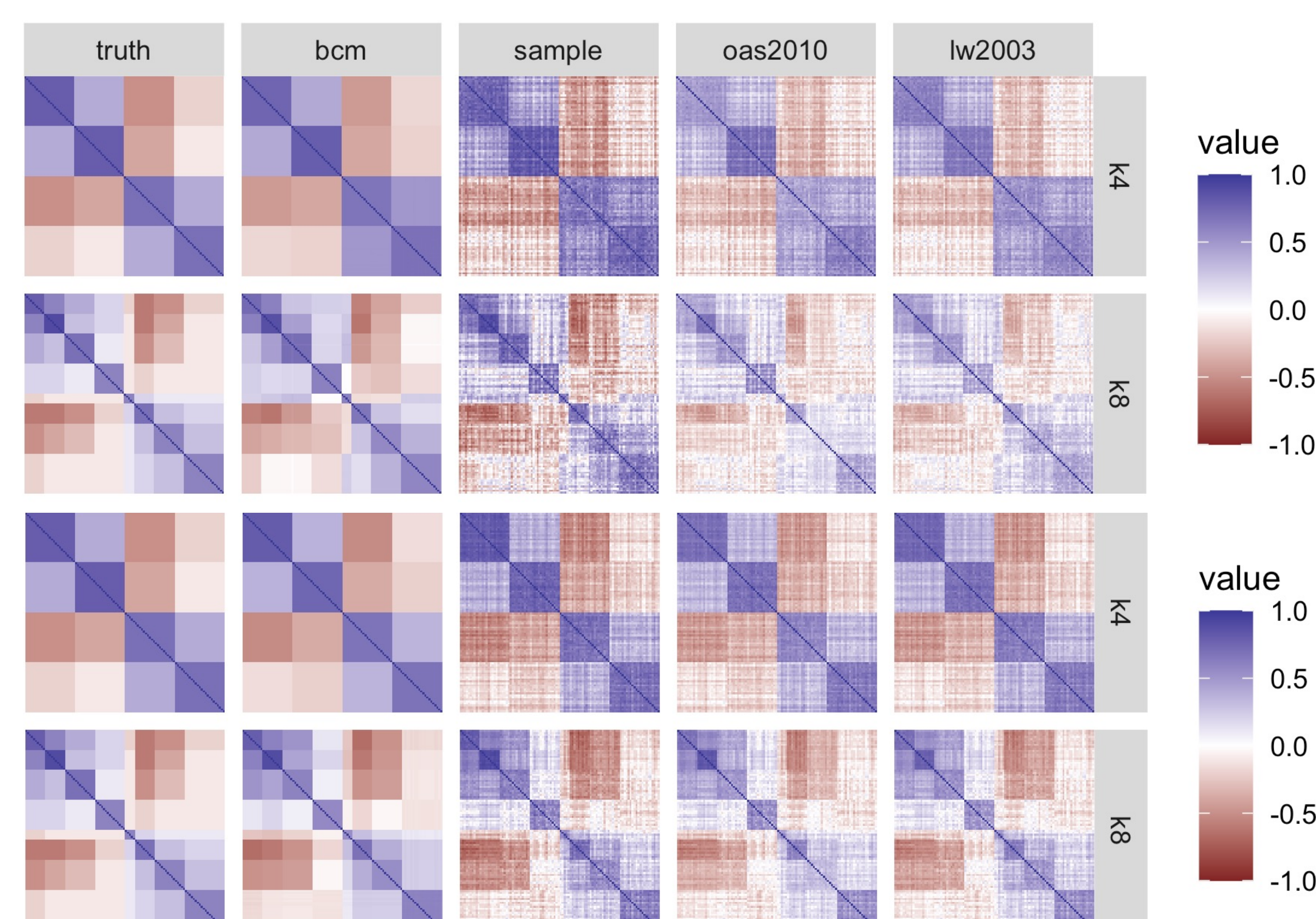


Fig. 2: Examples of induced priors for block correlation matrices under different group allocation

Numerical Experiments



- (n, p, k) : (# sample, # variables, # blocks)
- Estimation accuracy: BCM outperforms other alternatives except for situations with large p/k ratio
- Grouping: BCM recovers the true block structure in a decent way (smoothing/denoise) even under small n large p cases.

Fig. 3 (Left): Comparison between estimators when $p=100, n=20$ (up) or 50 (down), $k=4$ or 8

Fig. 4 (Right): Frobenius distance between estimators and truth under different (n, p, k) combinations

